

习题 4.4

1. $\frac{1}{9}$.

证明：由 Chebyshev 不等式，对任意的 $\varepsilon > 0$ ，有

$$P\left\{\left|\frac{1}{n}\sum_{i=1}^n X_i - \frac{1}{n}\sum_{i=1}^n EX_i\right| \geq \varepsilon\right\} \leq \frac{D\left(\frac{1}{n}\sum_{i=1}^n X_i\right)}{\varepsilon^2} = \frac{\frac{1}{n^2}D\left(\sum_{i=1}^n X_i\right)}{\varepsilon^2}$$

所以对任意的 $\varepsilon > 0$

$$\lim_{n \rightarrow \infty} P\left\{\left|\frac{1}{n}\sum_{i=1}^n X_i - \frac{1}{n}\sum_{i=1}^n EX_i\right| \geq \varepsilon\right\} \leq \frac{1}{\varepsilon^2} \lim_{n \rightarrow \infty} \frac{1}{n^2} D\left(\sum_{i=1}^n X_i\right) = 0$$

故 $\{X_n\}$ 服从大数定律。

2. $P\{|X - Y| \geq 6\} \leq \frac{1}{12}$. 3. 略. 4. 0.9842. 5. 142.

6. chebyshev 不等式: $n \geq 250$, 中心极限定理: $n = 68$.

解: 设 X 表示投掷一枚均匀硬币 n 次“正面向上”的次数，则

$$X \sim B(n, 0, 5), E(X) = np = 0.5n, D(X) = np(1-p) = 0.25n.$$

$$\begin{aligned} (1) P\left\{0.4 < \frac{X}{n} < 0.6\right\} &= P\{0.4n < X < 0.6n\} \\ &= P\{-0.1n < X - 0.5n < 0.1n\} \\ &= P\{|X - 0.5n| < 0.1n\} \\ &\geq 1 - \frac{D(X)}{(0.1n)^2} = 1 - \frac{0.25n}{0.01n^2} = 1 - \frac{25}{n} \geq 0.9, \end{aligned}$$

由此得 $\frac{25}{n} \leq 0.1, n \geq 250$.

$$\begin{aligned} (2) P\left\{0.4 < \frac{X}{n} < 0.6\right\} &= P\{0.4n < X < 0.6n\} \\ &= \Phi\left(\frac{0.6n - 0.5n}{\sqrt{0.25n}}\right) - \Phi\left(\frac{0.4n - 0.5n}{\sqrt{0.25n}}\right) \\ &= \Phi(0.2\sqrt{n}) - \Phi(-0.2\sqrt{n}) \\ &= 2\Phi(0.2\sqrt{n}) - 1 \geq 0.9, \end{aligned}$$

由此得

$$\Phi(0.2\sqrt{n}) \geq 0.95.$$

查表得 $0.2\sqrt{n} \geq 1.645, n \geq 67.65$, 取 $n = 68$.

习题四

一、填空题

1. 36.

2. $\frac{1}{2}$, 5.

分析：若 X 满足二项分布，则 $D(X) = np(1 - p)$,

$$\frac{dD(X)}{dp} = n(1 - p) - np = n(1 - 2p) = 0, p = \frac{1}{2},$$

$$\frac{d^2D(X)}{dp^2} \Big|_{p=\frac{1}{2}} = -2n < 0.$$

若 $p = \frac{1}{2}$ 是方差最大值点，也是标准差的最大值点，方差最大值为

$$np(1 - p) \Big|_{p=\frac{1}{2}} = 100 \times \frac{1}{2} \times \frac{1}{2} = 25,$$

从而标准差最大值为 $\sqrt{25} = 5$.

3. 1.

二、计算题

1. (1) $e^{-\frac{1}{8}} - e^{-\frac{1}{2}}$; (2) $\sqrt{2\pi}$.

2. (1) $\frac{1}{3}$; (2) $\frac{2}{\sqrt{\pi}}$.

3. (1) $a = 0.2$, $b = 0.1$, $c = 0.1$.

(2) Z 的概率分布为

Z	-2	-1	0	1	2
P	0.2	0.1	0.3	0.3	0.1

(3) $P\{X = Z\} = 0.4$.

4. $E(Y^2) = 5$.

解：因为

$$P\left\{X > \frac{\pi}{3}\right\} = \int_{\frac{\pi}{3}}^{\pi} \frac{1}{2} \cos \frac{x}{2} dx = \frac{1}{2},$$

所以 $Y \sim B\left(4, \frac{1}{2}\right)$ ，从而

$$EY = 4 \cdot \frac{1}{2} = 2, \quad DY = 4 \cdot \frac{1}{2} \cdot \frac{1}{2} = 1.$$

所以

$$EY^2 = DY + (EY)^2 = 1 + 2^2 = 5.$$

5. $E(X) = \frac{1}{p}$, $D(X) = \frac{1-p}{p^2}$.

6. $E(Z) = \frac{n^2 - 1}{3n}a$.

$$7. (1) Y \text{ 的分布列为 } P(Y = k) = C_{100}^k (0.6826)^k (0.3174)^{100-k};$$

$$(2) EY = 100 \times 0.6826 = 68.26, DY = 68.26 \times 0.3174 = 21.6657.$$

$$8. (1) E(X) = 0, D(X) = 2.$$

$$(2) \text{Cov}(X, |X|) = 0, X \text{ 与 } |X| \text{ 不相关};$$

$$(3) \text{ 随机变量 } X \text{ 与 } |X| \text{ 不独立.}$$

$$9. X \text{ 与 } Y \text{ 不相关不独立.}$$

解: 由于

$$E(X) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \sin \theta d\theta = 0, E(Y) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \cos \theta d\theta = 0.$$

$$\text{而 } E(XY) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \sin \cos \theta d\theta = 0. \text{ 因此}$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = 0,$$

从而 X 与 Y 不相关. 但由于 X 与 Y 满足关系

$$X^2 + Y^2 = 1,$$

所以 X 与 Y 不独立.

$$10. (1) X + Y \text{ 的概率分布.}$$

S	0	1	2	3
$P\{X + Y = S\}$	0.10	0.40	0.35	0.15

$$(2) 0.25;$$

$$(3) \rho_{XY} = -0.0676.$$

$$11. (1) E(Z) = \frac{1}{3}; D(Z) = 3;$$

$$(2) \rho_{XZ} = 0;$$

(3) X 与 Z 相互独立.

$$\text{解: (1)} E(Z) = E\left(\frac{X}{3} + \frac{Y}{2}\right) = \frac{1}{3}E(X) + \frac{1}{2}E(Y) = \frac{1}{3} \times 1 + \frac{1}{2} \times 0 = \frac{1}{3};$$

$$D(Z) = D\left(\frac{X}{3} + \frac{Y}{2}\right) = D\left(\frac{X}{3}\right) + D\left(\frac{Y}{2}\right) + 2\text{cov}\left(\frac{X}{3}, \frac{Y}{2}\right)$$

$$= \frac{1}{9}D(X) + \frac{1}{4}D(Y) + 2 \times \frac{1}{3} \times \frac{1}{2} \text{cov}(X, Y)$$

$$= \frac{1}{9} \times 3^2 + \frac{1}{4} \times 4^2 + \frac{1}{3} \cdot \rho_{XY} \cdot \sqrt{D(X) \cdot D(Y)}$$

$$= 1 + 4 + \frac{1}{3} \times \left(-\frac{1}{2}\right) \times 3 \times 4 = 3$$

$$(2) \rho_{XZ} = \frac{\text{cov}(X, Z)}{\sqrt{D(X) \cdot D(Z)}} = \frac{\frac{1}{2}[D(X+Z) - D(X) - D(Z)]}{\sqrt{D(X) \cdot D(Z)}}$$

$$\text{而 } D(X+Z) = D\left(X + \frac{X}{3} + \frac{Y}{2}\right) = D\left(\frac{4}{3}X + \frac{1}{2}Y\right)$$

$$\begin{aligned}
&= \left(\frac{4}{3}\right)^2 D(X) + \frac{1}{4}D(Y) + 2 \times \frac{4}{3} \times \frac{1}{2} \text{cov}(X, Y) \\
&= \frac{16}{9} \times 3^2 + \frac{1}{4} \times 4^2 + \frac{4}{3} \cdot \left(-\frac{1}{2}\right) \times 3 \times 4 = 12
\end{aligned}$$

$$\therefore \rho_{xz} = \frac{\frac{1}{2}(12 - 3^2 - 3)}{3 \times 3} = 0$$

(3) 由于 (X, Y) 服从二维正态分布, 且 $Z = \frac{X}{3} + \frac{Y}{2}$,

因此 (X, Z) 的联合分布也是二维正态分布, 根据 $\rho_{xz} = 0$ 可以确定 X 与 Z 相互独立.

12. (1) 常数 $a = \frac{1}{4}$, $b = -\frac{1}{4}$, $c = 1$;

(2) 方差 $D(X) = \frac{2}{3}$.

(3) $E(Y) = \frac{1}{4}(e^2 - 1)^2$, $D(Y) = \frac{1}{4}e^2(e^2 - 1)^2$.

13. (1)

	X_1		
X_2		0	1
0		0.1	0.8
1		0.1	0

(2) $P = -\frac{2}{3}$.

14. $\frac{\sigma_2^2 - \rho\sigma_1\sigma_2}{\sigma_1^2 + \sigma_2^2 - 2\rho\sigma_1\sigma_2}X_1 + \frac{\sigma_1^2 - \rho\sigma_1\sigma_2}{\sigma_1^2 + \sigma_2^2 - 2\rho\sigma_1\sigma_2}X_2$.

15. $E(X) = \frac{7}{6}$, $E(Y) = \frac{7}{6}$, $\text{Cov}(X, Y) = -\frac{1}{36}$, $\rho_{xy} = -\frac{1}{11}$, $D(X + Y) = \frac{5}{9}$.

16. (1) $f_V(v) = \begin{cases} 2e^{-2v}, & v > 0 \\ 0, & v \leq 0 \end{cases}$; (2) $E(U + V) = 2$.

解: (1) 设 V 的分布函数为 $F_V(v)$, 则

$$F_V(v) = P\{V \leq v\} = 1 - P\{V > v\} = 1 - P\{X > v\}P\{Y > v\}$$

当 $v < 0$ 时, $F_V(v) = 0$;

$$\text{当 } v \geq 0 \text{ 时, } F_V(v) = 1 - \int_v^{+\infty} e^{-x} dx \int_v^{+\infty} e^{-x} dx = 1 - e^{-2v}$$

故 V 的概率密度 $f_V(v) = F'_V(v) = \begin{cases} 2e^{-2v}, & v > 0 \\ 0, & v \leq 0 \end{cases}$

(2) 解法一: $U = \max\{X, Y\} = \frac{1}{2}[(X + Y) + |X - Y|]$

$$V = \min\{X, Y\} = \frac{1}{2}[(X + Y) - |X - Y|]$$

则 $U + V = X + Y$ $E(U + V) = E(X + Y) = E(X) + E(Y) = 1 + 1 = 2$.

解法二：同理， U 的概率密度 $f_U(u) = F'_U(u) = \begin{cases} 2(1 - e^{-u})e^{-u}, & u > 0 \\ 0, & u \leq 0 \end{cases}$

$$E(U + V) = E(U) + E(V) = \int_0^{+\infty} u 2(1 - e^{-u})e^{-u} du + \int_0^{+\infty} v 2e^{-2v} dv = \frac{3}{2} + \frac{1}{2} = 2.$$

17. (1) 0.8944; (2) 0.1379.

解： $X_i = \begin{cases} 1, & \text{第 } i \text{ 人治愈, } i = 1, 2, \dots, 100. \\ 0, & \text{其他.} \end{cases}$

$$\text{令 } x = \sum_{i=1}^{100} x_i.$$

(1) $X \sim B(100, 0.8)$,

$$\begin{aligned} P\left\{\sum_{i=1}^{100} x_i > 75\right\} &= 1 - P(X \leq 75) = 1 - \Phi\left(\frac{75 - 100 \times 0.8}{\sqrt{100 \times 0.8 \times 0.2}}\right) \\ &= 1 - \Phi(-1.25) = \Phi(1.25) = 0.8944. \end{aligned}$$

(2) $X \sim B(100, 0.7)$,

$$\begin{aligned} P\left\{\sum_{i=1}^{100} x_i > 75\right\} &= 1 - P(X \leq 75) = 1 - \Phi\left(\frac{75 - 100 \times 0.7}{\sqrt{100 \times 0.7 \times 0.3}}\right) \\ &= 1 - \Phi\left(\frac{5}{\sqrt{21}}\right) = 1 - \Phi(1.09) = 0.1379. \end{aligned}$$

18. $P\{X = k\} = C_{100}^k 0.2^k 0.8^{100-k}$, $k = 1, 2, \dots, 100$; 0.927.

解：(1) X 可看作 100 次重复独立试验中，被盗户数出现的次数，而在每次试验中被盗户出现的概率是 0.2，因此 $X \sim B(100, 0.2)$ ，故 X 的概率分布是

$$P\{X = k\} = C_{100}^k 0.2^k 0.8^{100-k}, k = 1, 2, \dots, 100.$$

被盗索赔户不少于 14 户且不多于 30 户的概率即为事件 $\{14 \leq X \leq 30\}$ 的概率。由中心极限定理，得

$$\begin{aligned} P\{14 \leq X \leq 30\} &= \Phi\left(\frac{30 - 100 \times 0.2}{\sqrt{100 \times 0.2 \times 0.8}}\right) - \Phi\left(\frac{14 - 100 \times 0.2}{\sqrt{100 \times 0.2 \times 0.8}}\right) \\ &= \Phi(2.5) - \Phi(-1.5) = 0.994 - [-0.33] = 0.927. \end{aligned}$$

19. $n \geq 14$.

解：设 X 表示要求使用外线的分机数，则 $X \sim b(200, 0.05)$ ，设总机处有 n 条外线，则

$$P\{X \leq n\} \geq 0.9,$$

由 De Moivre – Laplace 定理： $E(X) = np = 10$,

$$D(X) = np(1-p) = 9.5. P\{X \leq n\} = P\left\{\frac{X - 10}{\sqrt{9.5}} \leq \frac{n - 10}{\sqrt{9.5}}\right\} \approx \Phi\left(\frac{n - 10}{\sqrt{9.5}}\right) \geq 0.9 \text{ 查表}$$

$$\text{得 } \frac{n - 10}{\sqrt{9.5}} \geq 1.285 \quad n \geq 14.$$

20. 0.1814.

21. (1) 0.18; (2) 443.